Online Appendix Table 1
Adjusted and Unadjusted OLS Regression Coefficients for Fertility Outcomes

|  | Number of Pregnancies |  | Number of Births |  | Number of Living Children |  | Contraceptive Use |  | Pregnancy Termination |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Variable | Coef. | Std. <br> Err. | Coef. | Std. Err. | Coef. | Std. <br> Err. | Coef. | Std. <br> Err. | Coef. | Std. <br> Err. |
| $\hat{\tau}^{O L S}$ (Unadjusted) | 0.202 | 0.033 | 0.181 | 0.026 | 0.196 | 0.022 | -0.014 | 0.006 | 0.003 | 0.005 |
| $\hat{\tau}^{O L S}$ (Adjusted) | 0.204 | 0.023 | 0.189 | 0.017 | 0.184 | 0.015 | -0.016 | 0.005 | -0.001 | 0.005 |
| Number of Observations | 25366 |  | 25366 |  | 25366 |  | 25366 |  | 25366 |  |
| $R^{2}$ | 0.519 |  | 0.573 |  | 0.538 |  | 0.134 |  | 0.116 |  |

Note: This table compares the OLS estimates of a first-born girl on fertility outcomes with and without adjusting for the family level covariates. The first row reports the OLS regression of the fertility outcome on the first-born female dummy without additional covariates. The second row reports the OLS regression of the fertility outcome on the first-born female dummy after controlling for the first born's survival, year of survey, region, year of survey and region interactions, mother's age, age at first birth, years of education, ethnicity, rural residence, husband's age and years of education, patrilocal residence, whether the marriage was arranged and husband's family paid a bride price plus indicator variables for missing husbands age, husbands years of education, arranged marriage and bride price payment. Standard errors are heteroskedasticity-consistent.

Online Appendix Table 2
Coefficients for the Logit Regression

|  |  |  |  |  |
| :--- | ---: | :---: | :--- | :--- |
| Variable | Coefficient | Standard Error | $95 \%$ Confidence Interval |  |
|  |  |  |  |  |
| Mother's age | 0.0014 | 0.0029 | -0.0042 | 0.0070 |
| Mother's age at $1^{\text {st }}$ birth | -0.0067 | 0.0037 | -0.0139 | 0.0005 |
| Mother's years of education | 0.0065 | 0.0047 | -0.0027 | 0.0157 |
| Mother Non-Turkish | -0.0629 | 0.0404 | -0.1421 | 0.0162 |
| West | -0.0539 | 0.0418 | -0.1359 | 0.0280 |
| South | -0.0127 | 0.0448 | -0.1005 | 0.0750 |
| Central | -0.0597 | 0.0439 | -0.1458 | 0.0263 |
| North | -0.0876 | 0.0492 | -0.1840 | 0.0088 |
| Rural | -0.0128 | 0.0294 | -0.0704 | 0.0447 |
| Patrilocal Family | 0.0808 | 0.0434 | -0.0043 | 0.1659 |
| Father's age | 0.0015 | 0.0026 | -0.0036 | 0.0066 |
| Father's age missing | -0.0167 | 0.0477 | -0.1103 | 0.0769 |
| Father's years of education | 0.0013 | 0.0042 | -0.0068 | 0.0095 |
| Father's education missing | 0.0749 | 0.2041 | -0.3251 | 0.4750 |
| Arranged marriage | -0.0248 | 0.0280 | -0.0798 | 0.0301 |
| Arranged marriage missing | -0.4056 | 0.6322 | -1.6447 | 0.8334 |
| Bride price paid | 0.0432 | 0.0341 | -0.0237 | 0.1101 |
| Bride price payment missing | -0.0906 | 0.1053 | -0.2970 | 0.1157 |
| Survey year=1998 | 0.0151 | 0.0393 | -0.0620 | 0.0922 |
| Survey year=2003 | 0.0158 | 0.0378 | -0.0583 | 0.0899 |
| Survey year=2008 | -0.0106 | 0.0389 | -0.0868 | 0.0657 |
| Constant | -0.0433 | 0.1090 | -0.2569 | 0.1702 |
|  |  |  |  |  |
|  |  | Pseudo- $R^{2}=0.0006$ |  |  |

Note: This table reports the full set of coefficients from the logit regression of the first child's gender (equals 0 if a boy and 1 if a girl) on all the variables in the table. The first column reports the coefficients, the second column reports the standard errors of the coefficients, and the last two columns report the 95 percent confidence intervals for the estimated coefficients. The joint $\chi^{2}$-test results at the bottom are based on the null hypothesis that all the slope coefficients are jointly equal to zero.

## Online Appendix

Table 3. Interaction Effects on Family Size
Women Aged 15-49

| Survey Year <br> (1) |  |  |  | Mother's Education <br> (2) |  |  |  | Father's Education <br> (3) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ | Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ | Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ |
| 1993 | $\begin{aligned} & 0.162^{* * *} \\ & (0.031) \end{aligned}$ | 2.86 | $\begin{aligned} & 0.057 * * * \\ & (0.011) \end{aligned}$ | No Education | $\begin{aligned} & 0.255 * * * \\ & (0.046) \end{aligned}$ | 4.19 | $\begin{aligned} & 0.057 * * * \\ & (0.010) \end{aligned}$ | No Education | $\begin{gathered} 0.236 * * \\ (0.099) \end{gathered}$ | 4.73 | $\begin{aligned} & 0.049 * * \\ & (0.019) \end{aligned}$ |
| 1998 | $\begin{aligned} & 0.151^{* * *} \\ & (0.032) \end{aligned}$ | 2.79 | $\begin{aligned} & 0.054 * * * \\ & (0.011) \end{aligned}$ | Primary | $\begin{aligned} & 0.206 * * * \\ & (0.018) \end{aligned}$ | 2.51 | $\begin{aligned} & 0.082 * * * \\ & (0.007) \end{aligned}$ | Primary | $\begin{aligned} & 0.212 * * * \\ & (0.023) \end{aligned}$ | 2.97 | $\begin{aligned} & 0.069 * * * \\ & (0.007) \end{aligned}$ |
| 2003 | $\begin{aligned} & 0.211 * * * \\ & (0.028) \end{aligned}$ | 2.68 | $\begin{aligned} & 0.075 * * * \\ & (0.010) \end{aligned}$ | Secondary $\geq$ | $\begin{aligned} & 0.060 * * * \\ & (0.020) \end{aligned}$ | 1.79 | $\begin{aligned} & 0.036 * * * \\ & (0.010) \end{aligned}$ | Secondary $\geq$ | $\begin{aligned} & 0.143 * * * \\ & (0.017) \end{aligned}$ | 2.17 | $\begin{aligned} & 0.067 \text { *** } \\ & (0.007) \end{aligned}$ |
| 2008 | $\begin{aligned} & 0.203 * * * \\ & (0.028) \end{aligned}$ | 2.62 | $\begin{aligned} & 0.076 * * * \\ & (0.010) \end{aligned}$ |  |  |  |  |  |  |  |  |
| $p\left(\right.$ joint $\chi^{2}$ ) | 0.40 |  | 0.27 | $p\left(\right.$ joint $\chi^{2}$ ) | $<0.001$ |  | $<0.001$ | $p\left(\right.$ joint $\chi^{2}$ ) | 0.05 |  | 0.62 |
| $N$ | 25366 |  | 25366 |  | 25366 |  | 25366 |  | 25283 |  | 25283 |
| $R^{2}$ | 0.55 |  |  |  | 0.55 |  |  |  | 0.55 |  |  |

Note: This table shows the effect of a first-born female on sibship size for different subgroups estimated by interacting the first-born female dummy with each category of interest. The sample includes 25,366 women aged $15-49$. The outcome is the number of living children in the family. Panel (1) reports the effect of a first-born female on total number of living children by survey year. Panel (2) reports the effect of a first-born female on total number of living children by mother's education level. Panel (3) reports the effect of a first-born female on total number of living children by father's education level. The first column in each panel is estimated with OLS. The second column in each panel reports the mean number of children for families with first-born males, indicated with $\bar{y} \mid Z_{i}=0$. The third column in each panel is estimated with maximum likelihood assuming a Poisson process and shows the relative change in family size, $\% \Delta$, induced by a first-born female. The reported $p$-values are from $\chi^{2}$-tests based on the null hypothesis that the estimated coefficients are the same across the subgroups. All regressions control for the first born's survival, year of survey, mother's age, age at first birth, education level, ethnicity, region, rural residence, husband's age, husband's education level, patrilocal residence, whether the marriage was arranged and husband's family paid a bride price plus indicator variables for missing husband's age, husband's education, arranged marriage and bride price payment. Heteroskedasticity-consistent standard errors are in parentheses. Significance levels are indicated by $*<.10,{ }^{* *}<.05, * * *<.01$.

Online Appendix
Table 4. Interaction Effects on Family Size
Women Aged 35-49

| Survey Year <br> (1) |  |  |  | Mother's Education <br> (2) |  |  |  | Father's Education <br> (3) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ | Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ | Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ |
| 1993 | $\begin{aligned} & 0.193 * * * \\ & (0.058) \end{aligned}$ | 3.74 | $\begin{aligned} & 0.052 * * * \\ & (0.015) \end{aligned}$ | No Education | $\begin{aligned} & 0.239 * * * \\ & (0.060) \end{aligned}$ | 4.83 | $\begin{aligned} & 0.049 * * * \\ & (0.012) \end{aligned}$ | No Education | $\begin{gathered} 0.104 \\ (0.124) \end{gathered}$ | 5.42 | $\begin{gathered} 0.026 \\ (0.022) \end{gathered}$ |
| 1998 | $\begin{aligned} & 0.195 * * * \\ & (0.055) \end{aligned}$ | 3.58 | $\begin{aligned} & 0.056 * * * \\ & (0.015) \end{aligned}$ | Primary | $\begin{aligned} & 0.288^{* * *} \\ & (0.031) \end{aligned}$ | 3.03 | $\begin{aligned} & 0.090^{* * *} \\ & (0.009) \end{aligned}$ | Primary | $\begin{aligned} & 0.272 * * * \\ & (0.036) \end{aligned}$ | 3.60 | $\begin{aligned} & 0.071 * * * \\ & (0.009) \end{aligned}$ |
| 2003 | $\begin{aligned} & 0.270 * * * \\ & (0.045) \end{aligned}$ | 3.28 | $\begin{aligned} & 0.076 * * * \\ & (0.013) \end{aligned}$ | Secondary $\geq$ | $\begin{gathered} 0.085 * * \\ (0.036) \end{gathered}$ | 2.13 | $\begin{aligned} & 0.041^{* *} * \\ & (0.015) \end{aligned}$ | Secondary $\geq$ | $\begin{aligned} & 0.215^{* * *} \\ & (0.030) \end{aligned}$ | 2.66 | $\begin{aligned} & 0.078 * * * \\ & (0.010) \end{aligned}$ |
| 2008 | $\begin{aligned} & 0.264 * * * \\ & (0.044) \end{aligned}$ | 3.15 | $\begin{aligned} & 0.081 * * * \\ & (0.013) \end{aligned}$ |  |  |  |  |  |  |  |  |
| $p\left(\right.$ joint $\chi^{2}$ ) | 0.56 |  | 0.36 | $p\left(\right.$ joint $\chi^{2}$ ) | $<0.001$ |  | 0.004 | $p\left(\right.$ joint $\chi^{2}$ ) | 0.27 |  | 0.11 |
| $N$ | 12093 |  | 12093 |  | 12093 |  | 12093 |  | 12048 |  | 12048 |
| $R^{2}$ | 0.52 |  |  |  | 0.52 |  |  |  | 0.52 |  |  |

Note: This table shows the effect of a first-born female on sibship size for different subgroups estimated by interacting the first-born female dummy with each category of interest. The sample includes 12,093 women aged $35-49$. The outcome is the number of living children in the family. Panel (1) reports the effect of a first-born female on total number of living children by survey year. Panel (2) reports the effect of a first-born female on total number of living children by mother's education level. Panel (3) reports the effect of a first-born female on total number of living children by father's education level. The first column in each panel is estimated with OLS. The second column in each panel reports the mean number of children for families with first-born males, indicated with $\bar{y} \mid Z_{i}=0$. The third column in each panel is estimated with maximum likelihood assuming a Poisson process and shows the relative change in family size, $\% \Delta$, induced by a first-born female. The reported $p$-values are from $\chi^{2}$-tests based on the null hypothesis that the estimated coefficients are the same across the subgroups. All regressions control for the first born's survival, year of survey, mother's age, age at first birth, education level, ethnicity, region, rural residence, husband's age, husband's education level, patrilocal residence, whether the marriage was arranged and husband's family paid a bride price plus indicator variables for missing husband's age, husband's education, arranged marriage and bride price payment. Heteroskedasticity-consistent standard errors are in parentheses. Significance levels are indicated by $*<.10, * *<.05, * * *<.01$.

Online Appendix
Table 5. Interaction Effects on Family Size
Women Aged 15-49

| Patrilocal Residence <br> (1) |  |  |  | Arranged Marriage <br> (2) |  |  |  | Bride Price Paid <br> (3) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ | Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ | Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ |
| No | $\begin{aligned} & 0.185 * * * \\ & (0.016) \end{aligned}$ | 2.81 | $\begin{aligned} & 0.065 * * * \\ & (0.005) \end{aligned}$ | No | $\begin{aligned} & 0.147 * * * \\ & (0.021) \end{aligned}$ | 2.28 | $\begin{aligned} & 0.065 * * * \\ & (0.008) \end{aligned}$ | No | $\begin{aligned} & 0.163 * * * \\ & (0.015) \end{aligned}$ | 2.39 | $\begin{aligned} & 0.066 * * * \\ & (0.006) \end{aligned}$ |
| Yes | $\begin{aligned} & 0.179 * * * \\ & (0.038) \end{aligned}$ | 2.15 | $\begin{aligned} & 0.080^{* * *} \\ & (0.016) \end{aligned}$ | Yes | $\begin{aligned} & 0.208 * * * \\ & (0.020) \end{aligned}$ | 3.02 | $\begin{aligned} & 0.067 * * * \\ & (0.006) \end{aligned}$ | Yes | $\begin{aligned} & 0.250 * * * \\ & (0.042) \end{aligned}$ | 3.83 | $\begin{aligned} & 0.067 * * * \\ & (0.010) \end{aligned}$ |
| Difference | $\begin{aligned} & -0.006 \\ & (0.042) \end{aligned}$ |  | $\begin{gathered} 0.015 \\ (0.017) \end{gathered}$ | Difference | $\begin{gathered} 0.061 * * \\ (0.029) \end{gathered}$ |  | $\begin{gathered} 0.003 \\ (0.011) \end{gathered}$ | Difference | $\begin{gathered} 0.087 * \\ (0.044) \end{gathered}$ |  | $\begin{gathered} 0.001 \\ (0.012) \end{gathered}$ |
| $N$ | 25366 |  | 25366 |  | 25355 |  | 25355 |  | 24956 |  | 24956 |
| $R^{2}$ | 0.55 |  |  |  | 0.55 |  |  |  | 0.55 |  |  |

Note: This table shows the effect of a first-born female on sibship size for different subgroups estimated by interacting the first-born female dummy with each category of interest. The sample includes 25,366 women aged $15-49$. The outcome is the number of living children in the family. Panel (1) reports the effect of a first-born female on total number of living children by patrilocal residency. Panel (2) reports the effect of a first-born female on total number of living children by type of marriage. Panel (3) reports the effect of a first-born female on total number of living children by bride price payment to bride's family. The first column in each panel is estimated with OLS. The second column in each panel reports the mean number of children for families with first-born males, indicated with $\bar{y} \mid Z_{i}=0$. The third column in each panel is estimated with maximum likelihood assuming a Poisson process and shows the relative change in family size, $\% \Delta$, induced by a first-born female. The reported differences show if the estimated coefficients are the same across the two subgroups. All regressions control for the first born's survival, year of survey, mother's age, age at first birth, education level, ethnicity, region, rural residence, husband's age, husband's education level, patrilocal residence, whether the marriage was arranged and husband's family paid a bride price plus indicator variables for missing husband's age, husband's education, arranged marriage and bride price payment. Heteroskedasticity-consistent standard errors are in parentheses. Significance levels are indicated by $*<.10,{ }^{* *}<.05, * * *<.01$.

## Online Appendix

## Table 6. Interaction Effects on Family Size

Women Aged 35-49

| Patrilocal Residence <br> (1) |  |  |  | Arranged Marriage <br> (2) |  |  |  | Bride Price Paid <br> (3) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ | Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ | Category | OLS | $\bar{y} \mid Z_{i}=0$ | $\% \Delta$ |
| No | $\begin{aligned} & 0.233 * * * \\ & (0.025) \end{aligned}$ | 3.40 | $\begin{aligned} & 0.066 * * * \\ & (0.007) \end{aligned}$ | No | $\begin{aligned} & 0.230 * * * \\ & (0.039) \end{aligned}$ | 2.87 | $\begin{aligned} & 0.077 * * * \\ & (0.012) \end{aligned}$ | No | $\begin{aligned} & 0.216^{* * *} \\ & (0.025) \end{aligned}$ | 2.93 | $\begin{aligned} & 0.069 * * * \\ & (0.008) \end{aligned}$ |
| Yes | $\begin{aligned} & 0.321 * * \\ & (0.130) \end{aligned}$ | 3.49 | $\begin{aligned} & 0.090^{* *} \\ & (0.036) \end{aligned}$ | Yes | $\begin{aligned} & 0.238 * * * \\ & (0.032) \end{aligned}$ | 3.67 | $\begin{aligned} & 0.063 * * * \\ & (0.008) \end{aligned}$ | Yes | $\begin{aligned} & 0.281 * * * \\ & (0.062) \end{aligned}$ | 4.64 | $\begin{aligned} & 0.064 * * * \\ & (0.013) \end{aligned}$ |
| Difference | $\begin{aligned} & 0.088 \\ & (0.133) \end{aligned}$ |  | $\begin{gathered} 0.023 \\ (0.036) \end{gathered}$ | Difference | $\begin{gathered} 0.008 \\ (0.050) \end{gathered}$ |  | $\begin{aligned} & -0.014 \\ & (0.015) \end{aligned}$ | Difference | $\begin{gathered} 0.065 \\ (0.067) \end{gathered}$ |  | $\begin{aligned} & -0.004 \\ & (0.015) \end{aligned}$ |
| $N$ | 12093 |  | 12093 |  | 12087 |  | 12087 |  | 11916 |  | 11916 |
| $R^{2}$ | 0.52 |  |  |  | 0.52 |  |  |  | 0.52 |  |  |

Note: This table shows the effect of a first-born female on sibship size for different subgroups estimated by interacting the first-born female dummy with each category of interest. The sample includes 12,093 women aged $35-49$. The outcome is the number of living children in the family. Panel (1) reports the effect of a first-born female on total number of living children by patrilocal residency. Panel (2) reports the effect of a first-born female on total number of living children by type of marriage. Panel (3) reports the effect of a first-born female on total number of living children by bride price payment to bride's family. The first column in each panel is estimated with OLS. The second column in each panel reports the mean number of children for families with first-born males, indicated with $\bar{y} \mid Z_{i}=0$. The third column in each panel is estimated with maximum likelihood assuming a Poisson process and shows the relative change in family size, $\% \Delta$, induced by a first-born female. The reported differences show if the estimated coefficients are the same across the two subgroups. All regressions control for the first born's survival, year of survey, mother's age, age at first birth, education level, ethnicity, region, rural residence, husband's age, husband's education level, patrilocal residence, whether the marriage was arranged and husband's family paid a bride price plus indicator variables for missing husband's age, husband's education, arranged marriage and bride price payment. Heteroskedasticity-consistent standard errors are in parentheses. Significance levels are indicated by $*<.10, * *<.05, * * *<.01$.

Online Appendix
Table 7. Infant Mortality Differences among Third-born Children by Sex Composition of the Older Siblings

|  | Sex composition of the first two siblings |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) <br> Two Boys | (2) <br> One Boy \& One Girl |  | (3) <br> Two Girls |  |
|  | Third-born | Third-Born |  | Third-born |  |
|  | Boy | Boy | Girl | Boy | Girl |
| Mean <br> Standard Deviation | $\begin{aligned} & 0.081 \\ & {[0.27]} \end{aligned}$ | $\begin{aligned} & 0.068 \\ & {[0.25]} \end{aligned}$ | $\begin{aligned} & 0.068 \\ & {[0.25]} \end{aligned}$ | $\begin{aligned} & 0.060 \\ & {[0.24]} \end{aligned}$ | $\begin{aligned} & 0.071 \\ & {[0.26]} \end{aligned}$ |
| Girl-Boy difference |  |  |  |  |  |
| Difference-in-differences |  |  |  |  |  |
| Covariate Adjusted Difference-in-differences |  |  |  |  |  |
| $N$ |  |  |  |  |  |

Note: This table compares the infant mortality rates of the third-born children by sex composition of the older two siblings. Infant mortality is defined as the death of a child under the age of one. The sample is restricted to children who were born at least 12 months before the time of the interview. Girl-boy difference estimator shows the gender difference in infant mortality for third-born children by sex composition of the older two siblings. In panel (2), difference-in-difference estimator shows the difference in girl-boy infant mortality gap between third-born children who have no older male siblings and who have one female older sibling. In panel (3), difference-in-difference estimator shows the difference in girl-boy infant mortality gap between third-born children who have no older male siblings and who have no older female siblings. The covariate adjusted results are from the regressions that control for the first- and second-born older sibling's sex, year of survey, region, year of survey and region interactions, mother's age, age at first birth, years of education, ethnicity, rural residence, husband's age and years of education, patrilocal residence, whether the marriage was arranged and bride's family received a bride price plus indicator variables for missing husband's age, husband's years of education, arranged marriage and bride price payment. Heteroskedasticity-consistent standard errors are in parentheses. Significance levels are indicated by $*<.10, * *<.05, * * *<.01$.

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Table 8. The Effect of Previous Sibling's Sex on Child Vaccination

| Previous sibling's sex | $\begin{gathered} \text { (1) } \\ \text { BCG } \end{gathered}$ |  |  |  | $\begin{gathered} (2) \\ \text { DPT } \end{gathered}$ |  |  |  | (3) Polio |  |  |  | (4) MMR |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Boy |  | Girl |  | Boy |  | Girl |  | Boy |  | Girl |  | Boy |  | Girl |  |
| Subsequent sibling's sex | Boy | Girl | Boy | Girl | Boy | Girl | Boy | Girl | Boy | Girl | Boy | Girl | Boy | Girl | Boy | Girl |
| Outcome |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Mean | 0.814 | 0.816 | 0.817 | 0.810 | 0.819 | 0.820 | 0.833 | 0.817 | 0.873 | 0.872 | 0.884 | 0.866 | 0.690 | 0.680 | 0.699 | 0.685 |
| Standard Deviation | [0.39] | [0.39] | [0.39] | [0.39] | [0.38] | [0.38] | [0.37] | [0.39] | [0.33] | [0.33] | [0.32] | [0.34] | [0.46] | [0.47] | [0.46] | [0.46] |
| Girl-Boy difference | $\begin{gathered} 0.002 \\ (0.013) \end{gathered}$ |  | $\begin{aligned} & -0.007 \\ & (0.013) \end{aligned}$ |  | $\begin{gathered} 0.001 \\ (0.013) \end{gathered}$ |  | $\begin{aligned} & -0.016 \\ & (0.013) \end{aligned}$ |  | $\begin{aligned} & -0.000 \\ & (0.011) \end{aligned}$ |  | $\begin{aligned} & -0.018 \\ & (0.011) \end{aligned}$ |  | $\begin{aligned} & -0.010 \\ & (0.016) \end{aligned}$ |  | $\begin{aligned} & -0.013 \\ & (0.015) \end{aligned}$ |  |
| Difference-in-differences | $\begin{aligned} & -0.009 \\ & (0.018) \end{aligned}$ |  |  |  | $\begin{gathered} -0.017 \\ (0.018) \end{gathered}$ |  |  |  | $\begin{gathered} -0.018 \\ (0.016) \end{gathered}$ |  |  |  | $\begin{aligned} & -0.004 \\ & (0.022) \end{aligned}$ |  |  |  |
| Covariate adjusted |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Difference-in-differences | $\begin{aligned} & -0.007 \\ & (0.017) \end{aligned}$ |  |  |  | $\begin{aligned} & -0.017 \\ & (0.017) \end{aligned}$ |  |  |  | $\begin{aligned} & -0.017 \\ & (0.015) \end{aligned}$ |  |  |  | $\begin{gathered} 0.001 \\ (0.020) \end{gathered}$ |  |  |  |
| $N$ | 7,456 |  |  |  | 7,327 |  |  |  | 7,557 |  |  |  | 7,252 |  |  |  |

Note: This table compares the vaccination rates between boys and girls by previous sibling's sex. Regression samples are restricted to children who were born in the second birth parity or later and who were under the age of five at the time of the interview. Immunization outcomes are compared for BCG (Bacillus Calmette-Guerin), DPT (diphtheria, pertussis, tetanus), Polio, and MMR (measles-mumps-rubella) vaccinations in Panel (1) through Panel (4), respectively. Girl-boy difference estimator shows the gender difference in vaccination rates by previous sibling's sex. Difference-in-difference estimator shows the difference in girl-boy differences between children who has a previous female sibling and children who has a previous male sibling. The lower panel shows the same results from the regressions that control for the child's birth order, year of survey, region, year of survey and region interactions, mother's age, age at first birth, years of education, ethnicity, rural residence, husband's age and years of education, patrilocal residence, whether the marriage was arranged and husband paid a bride price plus indicator variables for missing husband's age, husband's years of education, arranged marriage and bride price payment. Standard errors are in parentheses and clustered by mother. Significance levels are indicated by $*<.10,{ }^{* *}<.05, * * *<.01$.

